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The chapter shows how to identify interactions that depend on the moderating variable using a SEM model. It uses the adaptive-gamma regression (AGRES) command in LISREL and how to interpret the results. In particular, we focus on the calculations of moderation and mediation effects in the chapter. To accomplish this the example in the book is extended and simplified so that it can be replicated and repeated. In this chapter, the authors explain and demonstrate several different types of

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heterogeneity. These include categorical heterogeneity, nominal heterogeneity, and subpopulation heterogeneity. The author then uses an example from the book to show how to specify each of these forms of heterogeneity using the LISREL software. They then illustrate how the relative fit of each model is compared using a likelihood ratio test. The chapter then ends by using an example to show how to conduct an F-test on two nested models. This chapter outlines the theory and computational algorithms involved in conducting a moderator effect analysis.

This is different from the approach described in the chapter on moderated moderation effect analysis. Unlike moderated moderation, the approach discussed in this chapter can be used with continuous or dichotomous moderators, and continuous outcomes. In order to demonstrate the use of the analytical procedure, we use an example

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from the book. The chapter concludes with a thorough discussion of the various moderation effects that have been proposed and discussed in the literature.

We also examine the problem of heterogeneous moderators and extend the analysis using the LISREL software.

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the freely estimated variance structure of the latent variable (measurement error) was contrasted to the model with equality constraints on the residual variances. the

fit of the freely estimated model was adequate (rmsea=.040, 90% ci (.007,.069), cfi=.986, rmsea=.042, 90% ci (.008,.076), cfi=.988), indicating that the latent scale was not significantly different from the observed variance.

hence, the test of the full metric invariance model was accepted. in the test of residual variances, the model with

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equal residual variances (rmsea=.049, 90% ci (.031,.068), cfi=.967, rmsea=.053, 90% ci (.034,.072), cfi=.965) did not fit the data as well as the freely estimated model (rmsea=.028, 90% ci (.002,.051), cfi=.992, rmsea=.031, 90% ci (.001,.993). the test of equality of residual variances was therefore rejected. hence, the model in which the variance of one of the factor loadings is equal to 1 was not a suitable model. in the test of scalar variance, the model with equality constraints on the factor loadings, intercepts, and residual variances (rmsea=.028, 90% ci (.002,.051), cfi=.992, rmsea=.031, 90% ci (.001,.993) did not fit the data well as the freely estimated model (rmsea=.053, 90% ci (.034,.072), cfi=.965). the test of equality of the intercepts was therefore rejected. hence, the model with free latent means and equal residual variances was not a suitable model. as

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the increasingly restrictive models estimate the same parameters as in the less constrained models, they are nested within the comparison model and have more degrees of freedom. hence, goodness of fit may be compared statistically by comparing whether the increase in chi-squares is significantly larger than the increase in degrees of freedom. as these tests were based on the rescaled satorra-bentler chi-square values, these difference tests were adjusted for non-normality according to instructions by satorra and bentler ( 2001 ). 5ec8ef588b

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